

# B.S.T.J. BRIEFS

## Axis-Crossing Intervals of Rayleigh Processes

By A. J. RAINAL

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### I. INTRODUCTION

Let  $R(t,a)$  denote the envelope of a stationary random process consisting of a sinusoidal signal of amplitude  $\sqrt{2a}$  and frequency  $f_0$  plus Gaussian noise of unit variance having a narrow-band power spectral density which is symmetrical about  $f_0$ . When  $a = 0$  Rice<sup>1</sup> presented some theoretical results which are very useful for studying statistical properties of the axis-crossing intervals of  $R(t,0)$ . The axis-crossing points and the axis-crossing intervals of the Rayleigh process  $R(t,a)$  are defined in Fig. 1. Some recent work concerning the axis-crossing intervals of  $R(t,a)$  was reported by Levin and Fomin<sup>2</sup>, Goryainov,<sup>3</sup> and Rainal.<sup>4,5</sup> The purpose of this brief is to present some theoretical results when  $a \geq 0$ . These results stem from a straightforward extension of Rice's analysis. The Rayleigh process  $R(t,a)$  occurs at the output of a typical radio or radar receiver during the reception of a sinusoidal signal immersed in Gaussian noise.

### II. THEORETICAL RESULTS

Using a notation consistent with Refs. 4 and 5 we define the following probability functions at an arbitrary level  $R$  of Fig. 1:

(1)  $Q^-(\tau, R, a)d\tau$ , the conditional probability that an upward axis-crossing occurs between  $t + \tau$  and  $t + \tau + d\tau$  given a downward axis-crossing at  $t$ .

(2)  $Q^+(\tau, R, a)d\tau$ , the conditional probability that a downward axis-crossing occurs between  $t + \tau$  and  $t + \tau + d\tau$  given an upward axis-crossing at  $t$ .

(3)  $[U(\tau, R, a) - Q(\tau, R, a)]d\tau$ , the conditional probability that an upward axis-crossing occurs between  $t + \tau$  and  $t + \tau + d\tau$  given an upward axis-crossing at  $t$ .

The reader should refer to Rice<sup>1</sup> for the definition of all notation which is not defined in this note. When  $a \geq 0$ , Rice's (86) becomes:

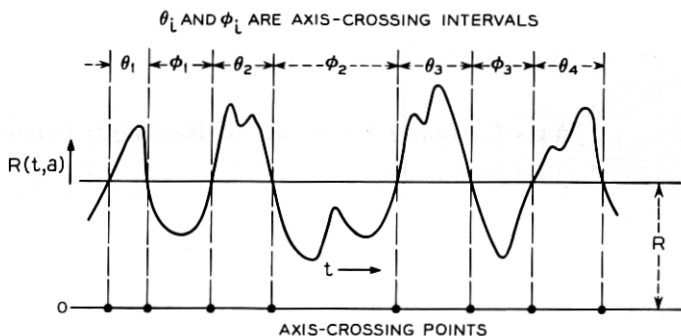


Fig. 1—The level  $R$  defines the axis-crossing points and the axis-crossing intervals of the Rayleigh process  $R(t, a)$ .

$$Q^-(\tau, R, a) = - \left( \frac{2\pi}{\beta} \right)^{\frac{1}{2}} R^{-1} e^{R^2/2} I_0^{-1}(RQ) e^a \quad (1)$$

$$\cdot \int_0^{2\pi} d\theta_1 \int_0^{2\pi} d\theta_2 \int_{-\infty}^0 dR_1' \int_0^{\infty} dR_2' R_1' R_2' p(R, R_1', R_2', R, \theta_1, \theta_2)$$

where:  $I_0(x)$  = Bessel function of imaginary argument

$$p(R, R_1', R_2', R, \theta_1, \theta_2) = \frac{R^2 [M_{22}^2 - M_{23}^2 c^2]^{-\frac{1}{2}}}{8\pi^3}$$

$$\cdot \exp \left\{ -\frac{1}{2M} [A(R_1'^2 + R_2'^2) + 2AR_1'R_2' + 2DR_1' + 2ER_2' + F] \right\}$$

$$A = [M_{22}^2 - M_{23}^2 c^2]^{-1} [MM_{22}(1 - m^2)] \quad Q = \sqrt{2a}$$

$$c = \cos(\theta_1 - \theta_2) \quad s = \sin(\theta_1 - \theta_2) \quad r = \frac{cM_{23}}{M_{22}}$$

$$D = [M_{22}^2 - M_{23}^2 c^2]^{-1} \{ R[M_{22} - M_{23}c][Mm'(c - m)]$$

$$+ Q[M_{12} - M_{13}][M_{23}s(M_{22} \sin \theta_2 + M_{23}c \sin \theta_1)$$

$$- \cos \theta_1 (M_{22}^2 - M_{23}^2 c^2)] \}$$

$$E = [M_{22}^2 - M_{23}^2 c^2]^{-1} \{ -R[M_{22} - M_{23}c][Mm'(c - m)]$$

$$+ Q[M_{12} - M_{13}][M_{23}s(M_{22} \sin \theta_1 + M_{23}c \sin \theta_2)$$

$$+ \cos \theta_2 (M_{22}^2 - M_{23}^2 c^2)] \}$$

$$F = [M_{22}^2 - M_{23}^2 c^2]^{-1} \{ 2[Q^2 - QR(\cos \theta_1 + \cos \theta_2)]$$

$$\cdot [M_{11} + M_{14}][M_{22}^2 - M_{23}^2 c^2] \}$$

$$\begin{aligned}
& - 2M_{13}QRs[M_{12} - M_{13}][M_{22} - M_{23}c][\sin \theta_1 - \sin \theta_2] \\
& - M_{22}Q^2[M_{12} - M_{13}]^2[\sin^2 \theta_1 + \sin^2 \theta_2] \\
& - 2M_{23}Q^2c[M_{12} - M_{13}]^2 \sin \theta_1 \sin \theta_2 \\
& + 2R^2[M_{22} - M_{23}c][(M_{22} + M_{23}c)(M_{11} + M_{14}c) - M_{13}^2s^2]\}.
\end{aligned}$$

Equation (1) can be put in a form analogous to Rice's (97) and (55):

$$Q^-(\tau, R, a) = \frac{e^a R M_{22} e^{R^2/2} I_0^{-1}(RQ)}{2\pi \sqrt{2\pi\beta} (1 - m^2)^2} \int_0^{2\pi} \int_0^{2\pi} e^{-(G/2M)} J(r, h, k) d\theta_1 d\theta_2 \quad (2)$$

where:

$$\begin{aligned}
J(r, h, k) & \equiv \frac{1}{2\pi s_1} \int_h^\infty dx \int_k^\infty dy (x - h)(y - k) e^z \\
z & = -\frac{x^2 + y^2 - 2rxy}{2(1 - r^2)}; \quad h = -a_1 \left[ \frac{1 - m^2}{M_{22}} \right]^{\frac{1}{2}}; \\
k & = a_2 \left[ \frac{1 - m^2}{M_{22}} \right]^{\frac{1}{2}}
\end{aligned}$$

$$a_1 = A^{-1}[1 - r^2]^{-1}[D - rE] \quad a_2 = A^{-1}[1 - r^2]^{-1}[E - rD]$$

$$G = A^{-1}[1 - r^2]^{-1}[2rDE - D^2 - E^2] + F; \quad s_1 = \sqrt{1 - r^2}.$$

We also find that:

$$\begin{aligned}
J(r, h, k) & = \frac{s_1}{2\pi} \exp \left[ -\frac{(k^2 - 2rhk + h^2)}{2(1 - r^2)} \right] - \frac{he^{-k^2/2}}{2\sqrt{2\pi}} \\
& \cdot \left[ 1 - P \left( \frac{h - rk}{s_1} \right) \right] - \frac{ke^{-h^2/2}}{2\sqrt{2\pi}} \left[ 1 - P \left( \frac{k - rh}{s_1} \right) \right] \\
& + (hk + r)K(r, h, k)
\end{aligned} \quad (3)$$

where:

$$P(x) = \frac{2}{\sqrt{2\pi}} \int_0^x e^{-t^2/2} dt$$

$$K(r, h, k) \equiv \text{Karl}^6 \text{ Pearson's } \left( \frac{d}{N} \right) = \frac{1}{2\pi s_1} \int_h^\infty dx \int_k^\infty dy e^z.$$

For a recent table of  $K(r, h, k)$  see Ref. 7. For a recent discussion of  $K(r, h, k)$  see the recent work of Gupta.<sup>8</sup> In these latter two references  $K(r, h, k)$  is denoted by  $L(h, k, r)$ .

$Q^+(\tau, R, a)$  is obtained from (1) by changing the signs of the  $\infty$ 's in

the limits of integration. We find that  $Q^+(\tau, R, a)$  is equal to the right-hand side of (2) with  $h, k$  replaced by  $-h, -k$ .

$[U(\tau, R, a) - Q(\tau, R, a)]$  is obtained from (1) by changing the lower limit of integration of  $R_1'$  to  $+\infty$ . We find that:

$$U(\tau, R, a) - Q(\tau, R, a) = \frac{e^a R M_{22} e^{R^2/2} I_0^{-1}(RQ)}{2\pi\sqrt{2\pi\beta} (1 - m^2)^2} \cdot \int_0^{2\pi} \int_0^{2\pi} e^{-(G/2M)} J_1(r, h, k) d\theta_1 d\theta_2 \quad (4)$$

where:

$$J_1(r, h, k) \equiv \frac{1}{2\pi s_1} \int_h^\infty dx \int_k^\infty dy (x - h) (y - k) e^x.$$

We find that  $J(r, h, k)$  and  $J_1(r, h, k)$  are related by:

$$J_1(r, h, k) = J(r, h, k) + \frac{h}{\sqrt{2\pi}} e^{-k^2/2} - \frac{(hk + r)}{2} [1 - P(k)]. \quad (5)$$

Equations (4) and (5) are the generalizations of (64) and (35) of Ref. 5.

### III. STATISTICAL DEPENDENCE OF AXIS-CROSSING INTERVALS

By expanding  $m(\tau)$  as:

$$m(\tau) = 1 - \frac{\beta}{2} \tau^2 + \frac{b_3 |\tau^3|}{3!} + \frac{b_4 \tau^4}{4!} + \frac{b_5 |\tau^5|}{5!} + \frac{b_6 \tau^6}{6!} + \frac{b_7 |\tau^7|}{7!} + o(\tau^7) \quad (6)$$

we find that as  $\tau \rightarrow 0$  from the right:

$$M_{11} = 2\beta b_3 \tau - (b_3^2 - \beta b_4 + \beta^3) \tau^2 + o(\tau^2) \quad (7)$$

$$M_{12} = \beta b_3 \tau^2 - \frac{1}{2}(b_3^2 - \beta b_4 + \beta^3) \tau^3 + o(\tau^3) \quad (8)$$

$$M_{13} = \beta b_3 \tau^2 - \frac{1}{2}(b_3^2 - \beta b_4 + \beta^3) \tau^3 + o(\tau^3) \quad (9)$$

$$M_{14} = -2\beta b_3 \tau + (b_3^2 - \beta b_4 + \beta^3) \tau^2 + o(\tau^2) \quad (10)$$

$$M_{22} = \frac{2}{3}\beta b_3 \tau^3 + \frac{1}{4}(\beta b_4 - b_3^2 - \beta^3) \tau^4 + o(\tau^4) \quad (11)$$

$$M_{23} = \frac{1}{3}\beta b_3 \tau^3 + \frac{1}{12}(3\beta b_4 - b_3^2 - 3\beta^3) \tau^4 + o(\tau^4). \quad (12)$$

When  $b_3 \neq 0$  we find that:

$$M_{12} - M_{13} = -\frac{1}{6}\beta^2 b_3 \tau^4 + o(\tau^4) \quad (13)$$

$$M_{11} + M_{14} = \frac{1}{6}\beta b_3^2 \tau^4 + o(\tau^4) \quad (14)$$

$$M_{22} - M_{23} = \frac{1}{3}\beta b_3 \tau^3 + o(\tau^3) \quad (15)$$

$$M = \frac{1}{3}\beta b_3^2 \tau^4 + o(\tau^4). \quad (16)$$

When  $b_3 = 0$  and  $b_5 \neq 0$  we find that:

$$M_{12} - M_{13} = \frac{1}{60}\beta^2 b_5 \tau^6 + o(\tau^6) \quad (17)$$

$$M_{11} + M_{14} = -\frac{1}{120}\beta b_4 b_5 \tau^7 + o(\tau^7) \quad (18)$$

$$M_{22} - M_{23} = -\frac{1}{30}\beta b_5 \tau^5 + o(\tau^5) \quad (19)$$

$$M = \frac{\beta b_5}{60} (\beta^2 - b_4) \tau^7 + o(\tau^7). \quad (20)$$

When  $b_3 = b_5 = 0$  we find that:

$$M_{12} - M_{13} = \frac{5\beta}{720} (\beta b_6 + b_4^2) \tau^7 + o(\tau^7) \quad (21)$$

$$M_{11} + M_{14} = \frac{-5b_4}{1440} (\beta b_6 + b_4^2) \tau^8 + o(\tau^8) \quad (22)$$

$$M_{22} - M_{23} = -\frac{1}{72} (\beta b_6 + b_4^2) \tau^6 + o(\tau^6) \quad (23)$$

$$M = \frac{1}{144} (\beta^2 - b_4) (\beta b_6 + b_4^2) \tau^8 + o(\tau^8). \quad (24)$$

As  $\tau \rightarrow 0$  we see that the terms of the quantities  $D$ ,  $E$ , and  $F$  which involve the sine wave amplitude  $Q$  are of higher order in  $\tau$  than the terms which do not involve  $Q$ . This behavior as  $\tau \rightarrow 0$  is consistent with a result reported by Levin and Fomin<sup>2</sup>. Thus, a theorem presented in Ref. 5 also applies to the Rayleigh process  $R(t, a)$ . That is: If  $R(t, a)$  is a Rayleigh process, defined in paragraph one, having a finite expected number of axis-crossing points per unit time at any level  $R$ , then two successive axis-crossing intervals at that level  $R$  are statistically dependent.

The theorem implies that the successive axis-crossing points of the Rayleigh process  $R(t, a)$  at any level  $R$  do not form a Markov point process.

#### IV. ACKNOWLEDGMENT

It gives me great pleasure to acknowledge stimulating discussions with S. O. Rice.

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8. Gupta, S. S., Probability Integrals of Multivariate Normal and Multivariate t, The Ann. of Math. Statistics, 34, No. 3, Sept. 1963, p. 792.

## Errata

A Note on a Signal Recovery Problem, by I. W. Sandberg, B.S.T.J., 43, November 1964, pp. 3065-3067.

On page 3066, replace  $|f_2(t)|^{\frac{1}{2}}$  by  $|f_2(t)|^2$ , and replace  $\tilde{\psi}[w] = \tilde{\psi}[w] - w$  by  $\tilde{\psi}[w] = \psi[w] - w$ . On page 3067, replace  $\max(c_1, c_2)$  by  $(c_1^2 + c_2^2)^{\frac{1}{2}}$ .